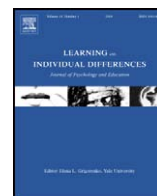




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# Raven's test performance of sub-Saharan Africans: Average performance, psychometric properties, and the Flynn Effect

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## ABSTRACT

This paper presents a systematic review of published data on the performance of sub-Saharan Africans on Raven's Progressive Matrices. The specific goals were to estimate the average level of performance, to study the Flynn Effect in African samples, and to examine the psychometric meaning of Raven's test scores as measures of general intelligence. Convergent validity of the Raven's tests is found to be relatively poor, although reliability and predictive validity are comparable to western samples. Factor analyses indicate that the Raven's tests are relatively weak indicators of general intelligence among Africans, and often measure additional factors, besides general intelligence. The degree to which Raven's scores of Africans reflect levels of general intelligence is unknown. Average IQ of Africans is approximately 80 when compared to US norms. Raven's scores among African adults have shown secular increases over the years. It is concluded that the Flynn Effect has yet to take hold in sub-Saharan Africa.

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## 1. Introduction

In 2007, Nobel laureate James Watson sparked controversy by expressing his gloom about the future of Africa in the light of the finding that sub-Saharan Africans have lower average IQ test scores than the peoples in the other parts of the world (The Sunday Times, October, 14, 2007). Although some scientists responded that his statements were “unsupported by science” (Federation of American Scientists, press release, October, 18, 2007), others responded more positively (Malloy, 2008; Rushton & Jensen, 2008). Notably, Lynn and Vanhanen (2002, 2006) collated the results of more than 50 published studies in which western IQ tests were administered to Africans. They concluded that the average IQ of the populations of sub-Saharan Africa in terms of UK norms lies below 70 (Lynn, 2006; Rushton & Jensen, 2005). Malloy (2008) also reviewed the literature on this issue and arrived at a similar conclusion. According to Rushton and Jensen (2009), this low average IQ suggests that African adults have the cognitive ability of an average 11-year-old white teenager. This assertion strikes us as unlikely and so the work by Lynn and Vanhanen and others raises the following questions: (1) Is their estimate of the average IQ of sub-Saharan Africans accurate? (2) How should we interpret IQ test performance of Africans? (3) Is IQ test performance of Africans a cause for concern about the prospects of sub-Saharan Africa? The aim of the present paper is to address these questions by presenting a balanced and critical evaluation of the present body of results of IQ testing in sub-Saharan Africa.

To estimate average IQ in countries all over the world, Lynn (and Vanhanen)<sup>1</sup> drew mainly on published data from Raven's Coloured Progressive Matrices (CPM; J. C. Raven, 1956) and the Standard Progressive Matrices (SPM; J. C. Raven, 1960). The Raven's tests are generally considered to be good non-verbal indicators of general intelligence or *g* (Carroll, 1993; Jensen, 1998), at least in western samples, and have often been administered in sub-Saharan Africa. For instance, Fahrmeier (1975) collected CPM data of schooled and unschooled Nigerian children. Lynn (and Vanhanen) compared their CPM scores to British norms<sup>2</sup> and calculated an average IQ of about 69 (Lynn, 2006; Lynn & Vanhanen, 2002). In another study conducted in Nigeria, Wober (1969) administered the SPM twice to a group of male factory workers. Lynn (and Vanhanen) compared their pretest scores to British norms, and concluded that their average IQ was below 65.

<sup>1</sup> Whenever we refer to “Lynn (and Vanhanen)” in this paper, we refer to Lynn and Vanhanen (2002, 2006) and to Lynn (2006), because the literature reviews in these three books overlap strongly. We refer to specific books whenever necessary.

<sup>2</sup> The estimation of IQ is described as follows: “Around 1973, data for the Coloured Progressive Matrices for a sample of 375 6–13 year-olds were collected by Fahrmeier (1975). In relation to the 1979 British standardization of the Standard Progressive Matrices, the mean IQ is 70. Because of the 6-year interval between the two data collections, this needs to be reduced to 69” (Lynn & Vanhanen, 2002, p. 215). Lynn (and Vanhanen) probably used a table provided on page 60 of the SPM manual (J. C. Raven et al., 1996) to convert raw CPM scores to raw SPM scores, to compare these CPM scores to British SPM norms of 1979. Note that their downward correction for outdated norms is an error because the norms are more recent than the test scores in Fahrmeier's sample. Hence, according to the appropriate use of this correction (i.e., 2 IQ points per decade), the IQ should have been raised by one point, not lowered by one.

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On the basis of these convenience samples, Lynn and Vanhanen (2002) claimed that the average IQ in Nigeria is below 70.

Lynn (and Vanhanen)'s work on IQ of Africans is often taken at face value (e.g., Herrnstein & Murray, 1994; Sarich & Miele, 2004; Skuy et al., 2002; Te Nijenhuis, De Jong, Evers, & van der Flier, 2004), even by their critics (e.g., MacEachern, 2006). The estimates of the average IQ of Africans are accorded a central role in the discussion on Black–White differences in IQ by Rushton and Jensen (2005, 2008, 2009), and these estimates feature in several evolutionary theories of intelligence (Kanazawa, 2004; Lynn, 2006; Rushton, 2000). Moreover, Lynn and Vanhanen's (2002, 2006) estimates of “national IQ” (i.e., an estimate of the average IQ of nations' inhabitants) have been used in at least twenty-five scientific studies (e.g., Barber, 2005; Dickerson, 2006; Gelade, 2008a,b; Jones & Schneider, 2006; Kanazawa, 2006, 2008; Ram, 2007; Rindermann, 2007, 2008a,b; Templer, 2008; Templer & Arikawa, 2006). Nevertheless, the work by Lynn (and Vanhanen) has also drawn criticism (Barnett & Williams, 2004; Ervik, 2003; Hunt & Carlson, 2007; Hunt & Sternberg, 2006). One point of critique is that Lynn and Vanhanen's estimate of average IQ among Africans is primarily based on convenience samples, and not on samples representative of the relevant populations (Barnett & Williams, 2004; Hunt & Sternberg, 2006). For example, the samples of Fahrmeier ( $N=375$ ) and Wober ( $N=86$ ) were neither intended, nor can be considered, to be representative of the entire population of Nigeria, a country of over 130 million inhabitants.

Lynn (and Vanhanen) did not take into account a sizeable portion of the relevant literature on IQ testing in sub-Saharan Africa. For instance, they did not consider other potentially relevant studies with the Raven's tests in Nigeria that indicated that average IQ in this country is considerably higher than 70 (Maqsud, 1980a,b; Okunrotifa, 1976). Perhaps the most important drawback of Lynn (and Vanhanen)'s reviews of the literature is that they are *unsystematic*. For example, nowhere in their reviews do they provide details concerning their literature search strategy, nor did they explicate the inclusion and exclusion criteria that they employed in selecting studies. It is well known that unsystematic literature reviews may lead to biased results (Cooper, 1998). The first aim of our review is to estimate the current level of average IQ of Africans on the basis of the Raven's tests by systematic literature review.

Because mean IQ levels in many western populations have shown marked increases over the course of the twentieth century (Flynn, 2006; Neisser, 1998), the interpretation of average IQ test scores of populations must be put into historical context. The scores on the Raven's tests have shown strong increases, a.k.a. Flynn Effects, in the developed world (Brouwers, van de Vijver, & van Hemert, 2009). Although there is some indication of a similar secular trend in IQs on the CPM in Kenya in recent years (Daley, Whaley, Sigman, Espinosa, & Neumann, 2003), little is known about the Flynn Effect in sub-Saharan Africa. The second aim of our review is therefore to determine whether there has been a Flynn Effect on the Raven's tests in sub-Saharan Africa. We note that the environmental variables that have been advanced as causes of the Flynn Effect in developed nations, such as improvements in nutrition and health (e.g., Lynn, 1990), and increases in educational opportunity and attainment (Ceci, 1991) are open to improvement in sub-Saharan Africa.

IQs do *not* necessarily equal a particular level of general intelligence or  $g$  (Bartholomew, 2004), as it is necessary to consider the issue of validity in interpreting an observed score as an indication of the position on a latent variable such as  $g$ . Several authors have questioned whether IQs of Africans are valid and comparable to scores in western samples in terms of general intelligence (Barnett & Williams, 2004; Ervik, 2003; Hunt & Carlson, 2007; Hunt & Sternberg, 2006). Some (e.g., Berry, 1974) reject the very possibility of obtaining a valid measure of  $g$  in sub-Saharan Africa with western IQ tests, while others (e.g., Herrnstein & Murray, 1994; Lynn, 2006; Rushton & Jensen, 2005) consider this relatively unproblematic. This issue is

related to our third question: what is the psychometric meaning of Africans' scores on the Raven's tests, and to what degree can we interpret these scores in terms of the cognitive abilities that these tests purport to measure? To this end, we consider studies that address the validity and reliability of these tests. We also consider the psychometric issue of measurement invariance (Mellenbergh, 1989; Millsap & Everson, 1993), which is crucial (e.g., Byrne et al., 2009) to the comparability of test scores across cultural groups in terms of latent variables, such as general intelligence.

Our paper is organized as follows. In Section 2, we present the results of a systematic review of the literature on the average performance of Africans on the basis of the Raven's tests. In Section 3, we consider whether there has been a Flynn Effect among Africans on these tests. Note that in Sections 2 and 3, we describe the mean performance and trends in the IQ metric. We do this to ensure consistency with the literature, but our use of the IQ metric should not be taken to imply that we consider scores on the Raven's tests of Africans valid indicators of general intelligence. In Section 4, we collate the results of studies that have addressed the validity the Raven's tests as measures of general intelligence or  $g$  in African samples. In this section, we view  $g$  in terms of Jensen's (1998) factor analytic definition (cf. Bartholomew, 2004), which is closely related to other views on intelligence (Carroll, 1993; McGrew, 2005). We do this because this definition is amendable to psychometric modeling, but we recognize that the interpretation of  $g$  remains open to debate (Bartholomew, Deary, & Lawn, 2009; van der Maas et al., 2006). Hence, Section 4 is not concerned with the nature of  $g$  itself, but rather provides an empirical test of the prediction from the  $g$  theory (Jensen, 1980, 1998) that Raven's tests are valid indicators of  $g$  among Africans (Lynn, 2006; Lynn & Vanhanen, 2006; Rushton & Jensen, 2005; Rushton & Jensen, 2009; Rushton & Skuy, 2000).

## 2. Is average IQ of Africans really below 70?

In a previous study (Wicherts, Dolan, & van der Maas, 2010), we considered the IQ test performance of Africans on several western IQ tests and found that their average IQ is appreciably higher than Lynn (and Vanhanen)'s estimates suggest.<sup>3</sup> In the current study, we attempted to locate all the published data of Africans on the Raven's tests. We explicitly discuss the criteria we used for inclusion and exclusion of particular studies. Our specific aims are to estimate the average IQ of African samples on the SPM and CPM. In Section 3, we determine whether a Flynn Effect is present in African populations. Our literature review also provided us with the opportunity to consider the psychometric properties of the Raven's tests in African samples, which we address in Section 4. Note that our review of the literature and our estimates of average IQ are concerned with the overall Black population of sub-Saharan Africa. This necessarily represents a crude generalization that can not do justice to the wide cultural, social, and economic variation in sub-Saharan Africa. We conduct more fine grained analyses whenever possible.

### 2.1. Method

#### 2.1.1. Selection bias

It is well known that convenience samples may produce biased results. In estimating average scores of the population of (countries in) sub-Saharan Africa, Lynn (and Vanhanen) used published studies, which included small convenience samples (e.g., Fahrmeier, 1975; Wober, 1966) and large representative samples (e.g., Costenbader &

<sup>3</sup> At a time when the current manuscript was ready for submission, reviewers of our other paper in *Intelligence* demanded that we include in that paper some data on the mean performance of Africans on the Raven's from the current review. We view both studies as replications of the same issue and report the main results in the current paper.

Ngari, 2001; MacArthur, Irvine, & Brimble, 1964). In estimating the average IQ of Africans, they assigned equal weight to both types of sample. Because in most cases representative samples are much larger than convenience samples, one straightforward, albeit partial, solution to the issue of selection bias is to weight average IQs by sample size. An additional reason to do so, is that the effect of sampling variability decreases as sample size increases. We include in our review both convenience and representative samples, but specifically consider the latter type of samples.

### 2.1.2. Search of studies

Besides considering the sources used by Lynn (and Vanhanen) and Malloy (2008), we located studies in which the SPM or CPM were administered in African samples. To this end, we submitted a combination of various search terms to PsycINFO. These included “Raven”, “SPM”, “CPM”, “progressive matrices”, “IQ”, “intelligence”, “cognitive ability”, “abilities” combined with the words “Africa”, “African”, and the names and adjectives of all countries in the continent (e.g., “Nigeria” or “Nigerian”). We also located additional papers in the reference lists of the papers that we found. In addition, we identified in Web of Science all articles (from 1988 to 2006) referring to the various SPM and CPM manuals. This resulted in about 2500 papers concerning the SPM/CPM. The titles and authors’ names of all these papers were scanned for relevance. We used only books, papers, or reports that were available through the IBL system in the Netherlands, a system connecting 400 Dutch libraries. Although our approach resulted in a large number of studies of the SPM and CPM in Africa, it is conceivable that we missed other studies.

### 2.1.3. Exclusion criteria

The following exclusion criteria were employed in the selection of studies. First, we excluded samples that are clearly abnormal, such as a sample of physically handicapped adults (Osuji, 1985). Furthermore, the condition of administration of the tests should reasonably approximate those stipulated in the test manual. For instance, we excluded the SPM scores of Zindi’s (1994) sample of Zimbabwean school children, because the SPM was not administered in its entirety (36 of the 60 items were administered), and because it is unclear how Zindi arrived at his IQ estimate of 70. We also disregarded Klingelhofer’s study using a large sample of secondary school students from Tanzania, because Klingelhofer imposed a time limit on the SPM (a non-standard condition). He did so explicitly to “[preclude] some of the kinds of comparisons that have marked the literature” (Klingelhofer, 1967, p. 206). Whenever tests were administered twice, we used the pretest scores. Although the potential unfamiliarity of some African test-takers with cognitive testing may render their scores on a retest more valid (Laroche, 1959; Wober, 1969), retest effects make it difficult to compare the mean results from retests to western norms that were based on a single test administration. We did not assign IQ values to studies in which the SPM or CPM did not meet basic psychometric standards, as will be discussed in the results section. For instance, the test–retest reliability in Wober’s sample of Nigerian factory workers was 0.59, i.e., below the 0.80 typically found with the SPM (J. C. Raven, Court, & Raven, 1996), and the correlation between pretest SPM scores and educational attainment did not deviate significantly from zero (see also Wober, 1966). Accordingly, we discarded this sample in estimating the average IQ on the Raven’s tests.

It is quite likely that African test-takers experience mundane problems that may lead to an underestimation of their latent cognitive ability, such as the misunderstanding of instructions (Irvine, 1966; Kendall, Verster, & Von Mollendorf, 1988), or suboptimal testing conditions (Sternberg, 2004). For instance, in the Fahrmeier’s study with the CPM in Nigeria, “children were tested on porches, in entrance rooms, or under trees” (Fahrmeier, 1975, p. 282) by untrained personnel. This runs counter to the official guidelines: “The person

to be tested is seated comfortably opposite the psychologist at a table about 2 feet wide” (J. C. Raven, 1956, p. 13, italics added). As non-standard test settings may depress performance, we excluded from our review the studies in which such effects were obvious. However, not all data sources included sufficient information to establish whether testing conditions were acceptable. We stress, therefore, the importance of exercising care in interpreting the IQs of Africans, which we provide below.

We only used data sets from sources, which included sufficient descriptive statistics. We thus excluded a large number of studies in which raw means, medians, or percentile scores were absent. We excluded CPM data of age ranges for which no British norms exist. This criterion resulted in the exclusion of several studies in which the CPM was administered to adolescents and adults (Berlioz, 1955; Berry, 1966; Binnie Dawson, 1984; Boissiere, Knight, & Sabot, 1985; Kendall, 1976; Sternberg et al., 2001). Lynn (and Vanhanen) and Malloy (2008) assigned average IQs below 70 to these samples (cf. Herrnstein & Murray, 1994). However, there are no (British) CPM norms above the age of 11. Lynn (and Vanhanen) (R. Lynn, personal communication, June 22, 2006) employed a table from the SPM manual (J. C. Raven et al., 1996) with which CPM scores can be converted to SPM scores (cf. Lynn, 1997). These approximate SPM scores can be compared to adult norms, allowing a rough estimate of IQ. However, this method does not result in accurate estimates of IQ, because the CPM is too easy for normal test-takers above the age of 11. The resulting ceiling effect makes it difficult to get an above-average SPM norms IQ on the CPM (e.g., for young adults, a score of 34 out of 36 corresponds to an IQ of 93). This virtually guarantees an underestimation of IQ with the CPM in samples above the age of 11, particularly in adults. The drawbacks of this conversion method are evidenced by the fact that it accords an average IQ of 75 to a carefully selected norm sample of 894 normal adults from Italy and San Marino (Measso et al., 1993). Because this conversion method does not result in reasonable IQ estimates, we did not consider mean IQs based on CPM scores of adults and adolescents. Appendix B (Section 6.2) provides an overview of all excluded samples (for which sufficient descriptive statistics were available), together with the reasons for exclusion. Obviously, the use of alternative inclusion criteria may result in different estimates of the average IQ. Below, we evaluate the effects of our choices to exclude samples on the ultimate estimate.

### 2.1.4. Converting raw scores to IQ

SPM and CPM raw scores need to be converted to percentile scores given in British norm tables. These percentile score are then transformed to standard IQs ( $M=100$  and  $SD=15$ ). Details of this conversion method and a discussion of its problems are given in Appendix A (Section 6.1). For the data collected up to 1965 (i.e., the midpoint of various SPM/CPM standardizations), we used older norm tables (J. C. Raven, 1956, 1960). For more recent samples (after 1965), we used recent norm tables (J. Raven, 2000; J. C. Raven, Court, & Raven, 1990; J. C. Raven et al., 1996). We provide IQ estimates with and without a Flynn Effect correction for outdated norms. This correction for the Flynn Effect is consistent with Lynn (and Vanhanen)’s correction.

## 2.2. Results

### 2.2.1. Raven’s Standard Progressive Matrices

Table 1 gives average SPM scores and corresponding IQs in 40 samples of Africans, totaling 15,408 cases. The table includes the country of origin, a short description of the sample, the sample size, the approximate or reported year of testing, the age range or average age, the percentage of formally schooled (i.e., more than 3 years of education) persons in each sample, the average raw score, our IQ estimates, and the IQ estimates provided by Lynn (2006). The average IQ per sample varies between 69 and 97 compared to an average IQ of 100 in Great Britain. Combining these averages results in an



**Table 1**  
Sub-Saharan African scores on the Standard Progressive Matrices.

| Source                                | Country             | Sample description   | N    | Year   | Age   | Edu | M    | IQ | IQ FE | IQ Lynn |
|---------------------------------------|---------------------|--|------|--------|-------|-----|------|----|-------|---------|
| Crawford Nutt (1976)                  | South Africa        | Children from high school in Soweto                                      | 228  | ± 1975 | 19    | 100 | 45.0 | 83 | 84    | –       |
| Fontaine (1963)                       | Mali                | Secondary school children  | 790  | 1962   | 14–17 | 100 | ≈26  | 79 | 76    | –       |
| Fontaine (1963)                       | Mali                | Professional adults  | 270  | 1962   | 17+   | 100 | ≈22  | 74 | 71    | –       |
| Grieve & Viljoen (2000), Sonke (2001) | South Africa        | Impoverished University students in Venda                                | 30   | 1996   | 19–29 | 100 | 37.4 | 75 | 74    | 77      |
| Ijarotimi & Ijadunola (2007)          | Nigeria             | Primary school children in Ondo State                                    | 402  | ± 2005 | 11.5  | 100 | 20.9 | 74 | 69    | –       |
| Irvine (1969b)                        | Zimbabwe            | Random selection of children with 8 yr of education                      | 200  | 1962   | 14–18 | 100 | 27.8 | 81 | 77    | –       |
| Jedege & Bamgboye (1981)              | Nigeria             | Random sample of secondary school students in Oyo State                  | 755  | 1977   | 11–15 | 100 | 28.5 | 77 | 77    | –       |
| Kaniel & Fisherman (1991)             | Ethiopia            | Uneducated Ethiopian Jews in Israel                                      | 250  | ± 1985 | 14–15 | 0   | 27   | 69 | 68    | 69      |
| Kozulin (1998)                        | Ethiopia            | Ethiopian Jews immigrated to Israel                                      | 46   | ± 1995 | 14–16 | 100 | 28.4 | 72 | 69    | 65      |
| Laroche (1959)                        | Dem. Rep. Congo     | Boys in schools in Elizabethville  | 222  | 1955   | 10–15 | 100 | 29.5 | 86 | 83    | 68      |
| Latouche & Dormeau (1956)             | Cent. Afr. Republic | Candidates for centre for accelerated technical learning in Bangui       | 1144 | ± 1953 | 17+   | 100 | 19.5 | 72 | 71    | 64      |
| Latouche & Dormeau (1956)             | Congo–Braz.         | Candidates for centre for accelerated technical learning in Brazzaville  | 1596 | ± 1953 | 17+   | 100 | 23.9 | 78 | 77    | 64      |
| Latouche & Dormeau (1956)             | Congo–Braz.         | Candidates for centre for accelerated technical learning in Pointe-Noire | 580  | ± 1953 | 17+   | 100 | 23.6 | 78 | 77    | –       |
| Lynn & Holmshaw (1990)                | South Africa        | Children from state primary schools                                      | 350  | 1988   | 9.5   | 100 | 12.7 | 77 | 75    | 63      |
| MacArthur et al. (1964)               | Zambia              | Repres. sample of class 6 students                                       | 759  | 1963   | 15.5  | 100 | ≈27  | 79 | 75    | 77      |
| MacArthur et al. (1964)               | Zambia              | Repres. sample of Form II students                                       | 649  | 1963   | 17.5  | 100 | ≈34  | 87 | 83    | –       |
| MacArthur et al. (1964)               | Zambia              | Technical college students in Lusaka                                     | 195  | 1963   | 18+   | 100 | ≈30  | 83 | 79    | –       |
| MacArthur et al. (1964)               | Zambia              | Mine farm youth students   | 292  | 1963   | 16.5  | 100 | ≈26  | 79 | 75    | –       |
| Maqsd (1997)                          | South Africa        | Batswana tribe high school students                                      | 140  | ± 1995 | 17–20 | 100 | ≈39  | 75 | 75    | –       |
| Maqsd (1980b)                         | Nigeria             | Secondary school girls in Kano city                                      | 136  | ± 1979 | 13–15 | 100 | 38.7 | 85 | 85    | –       |
| Maqsd (1980a)                         | Nigeria             | Primary school boys in Kano city   | 120  | ± 1979 | 11–12 | 100 | 22.1 | 72 | 72    | –       |
| Morakinyo (1985)                      | Nigeria             | Psychiatric out-patients and controls                                    | 28   | ± 1983 | 18+   | ?   | ≈47  | 87 | 86    | –       |
| Nkaya, Huteau, & Bonnet (1994)        | Congo–Braz.         | Secondary school children in Brazzaville                                 | 88   | ± 1992 | 13.25 | 100 | 29.6 | 75 | 72    | 73      |
| Notcutt (1950)                        | South Africa        | Zulus in primary schools near Durban                                     | 1008 | 1948   | 8–16  | 100 | 22.5 | 81 | 82    | 75      |
| Notcutt (1950)                        | South Africa        | Literate and illiterate Zulu adults                                      | 703  | 1949   | 17+   | 44  | 22.2 | 75 | 75    | 64      |
| Nwuga (1977)                          | Nigeria             | Children with a history of kwashiorkor and normal controls               | 165  | ± 1975 | 9–10  | 100 | 21.9 | 82 | 83    | –       |
| Okunrotifa (1976)                     | Nigeria             | Rural primary school children  | 50   | 1974   | 5.5   | 100 | ≈12  | 87 | 88    | –       |
| Okunrotifa (1976)                     | Nigeria             | Urban primary school children  | 100  | 1974   | 7.0   | 100 | ≈13  | 84 | 85    | –       |
| Ombredane, Robaye, & Robaye, (1957)   | Dem. Rep. Congo     | Members of Baluba tribe  | 320  | 1954   | 17–29 | 74  | 22.1 | 75 | 74    | 64      |
| Owen (1992)                           | South Africa        | Children from secondary schools  | 1093 | 1986   | 16    | 100 | 27.7 | 69 | 68    | 63      |
| Pons (1974)                           | Zambia              | Bemba males employed in mining   | 152  | ± 1961 | 18+   | 100 | 23.2 | 77 | 75    | 64      |
| Pons (1974)                           | Zambia              | Bemba males employed in mining   | 1011 | ± 1965 | 18+   | 100 | 33.7 | 87 | 84    | –       |
| Raveau et al. (1976)                  | Madagascar          | African adults working in France   | 143  | ± 1975 | 18+   | 100 | 40.9 | 79 | 82    | 82      |
| Raveau et al. (1976)                  | Various             | African adults working in France   | 588  | ± 1975 | 18–49 | 100 | 38.5 | 74 | 77    | –       |
| Rushton & Skuy (2000)                 | South Africa        | University students in psychology  | 173  | 1998   | 17–23 | 100 | 43.3 | 80 | 79    | 83      |
| Rushton et al. (2002)                 | South Africa        | University students in engineering                                       | 198  | ± 2000 | 17–23 | 100 | 50   | 92 | 90    | 93      |
| Skuy et al. (2002)                    | South Africa        | University students in psychology  | 70   | ± 2000 | 17–29 | 100 | 43.2 | 80 | 78    | 81      |
| Sonke (2001)                          | South Africa        | Illiterates from rural Venda   | 17   | 1995   | 13–20 | 50  | 25.7 | 69 | 66    | 68      |
| Taylor (2008)                         | South Africa        | Adult job applicants   | 200  | 2006   | 33.8  | ?   | 41.2 | 78 | 75    | –       |
| Zaaiman et al. (2001)                 | South Africa        | Disadvantaged university students  | 147  | 1995   | 18+   | 100 | 52.3 | 97 | 96    | 100     |

Note: M: raw mean SPM score; Edu: Percentage of sample formally educated; IQ: IQ estimate; IQ FE: IQ estimate with correction for Flynn Effect; IQ Lynn: IQ estimate by Lynn (2006); and ≈ indicates median values.

*N*-weighted mean IQ of 78.4 (median 78, *SD* = 5.6). Average IQ on the SPM of people in the United States is approximately 2 points lower than the UK average (Lynn & Vanhanen, 2002; J. C. Raven et al., 1996). If we choose to compare the average SPM scores of Africans to an IQ of 100 for the United States, average SPM IQ of Africans would be 80. A Flynn Effect correction would lower these estimates by about one IQ point.

The samples, considered by Lynn (and Vanhanen), but discarded here, are given in the Appendix. Besides the two samples described above (Klingelhofer, 1967; Zindi, 1994), these are Wober's (1969) sample of factory workers, and Verhaegen's (1956) sample of uneducated adults from a primitive tribe in the then Belgian Congo in the 1950s. Verhaegen indicated that the SPM test format was rather confusing to the test-takers, and that the test did not meet the standards of valid measurement. In Wober's study, the reliability and validity were too low (Wober, 1975). In three of the samples in Table 1, the average IQ is below 70. These are Owen's large sample of Black South African school children tested in the 1980s, the 17 Black South Africans carefully selected for their illiteracy by Sonke (2001), and a group of uneducated Ethiopian Jewish children, who lived

isolated from the western world in Ethiopia and immigrated to Israel in the 1980s (Kaniel & Fisherman, 1991). The last two samples cannot be considered to be representative.

Representative samples include Irvine's (1969b) random selection from the 1962 standardization sample of schooled children in Zimbabwe (then Southern Rhodesia), the standardization sample of the Northern Rhodesia Mental Survey (MacArthur et al., 1964), Notcutt's (1950) standardization samples of Zulu school children, and literate and illiterate Zulu adults in South Africa, and Jedege and Bamgboye's (1981) sample of secondary school students in Nigeria. These more carefully sampled groups of test-takers all show average IQs of 75 or higher. If we were to consider only these representative samples (Irvine, 1969b; MacArthur et al., 1964; Notcutt, 1950; Owen, 1992), the (*N*-weighted) average IQ would be 77.5.

Note that five of the SPM samples reviewed here contain Black university students from South Africa (Grieve & Viljoen, 2000; Rushton & Skuy, 2000; Rushton, Skuy, & Fridjhon, 2002; Skuy et al., 2002; Zaaiman, van de Vijver, & Thijs, 2001). These students (*N* = 618) score higher on average (IQ: *M* = 88) than the remaining samples. In contrast to the university students, 734 cases (4.8%) in Table 1 had no

**Table 2**  
Sub-Saharan African scores on the coloured Progressive Matrices.

| Source                     | Country         | Sample description  | N     | Year   | Age    | Edu | M     | IQ    | IQFE  | IQ Lynn |
|----------------------------|-----------------|---|-------|--------|--------|-----|-------|-------|-------|---------|
| Aboud et al. (1991)        | Ethiopia        | Children in an orphanage in Jimma                                 | 134   | ± 1989 | 5–11   | 100 | 13.56 | 72    | 71    | –       |
| Ayalew (2005)              | Ethiopia        | Rural children  | ± 108 | 1994   | 7–11   | ?   | 14.2  | 72    | 70    | –       |
| Costenbader & Ngari (2001) | Kenya           | Representative sample of school children                          | 1222  | ± 1998 | 6–10   | 100 | 15.86 | 82    | 79    | 75      |
| Daley et al. (2003)        | Kenya           | Children from rural district of Embu                              | 118   | 1984   | 7.5    | 100 | 12.82 | 75    | 75    | 76      |
| Daley et al. (2003)        | Kenya           | Children from rural district of Embu                              | 537   | 1998   | 7.5    | 100 | 17.31 | 90    | 87    | 89      |
| Fahrmeier (1975)           | Nigeria         | Schooled and unschooled children                                  | 334   | ± 1973 | 6–11   | 57  | 11.42 | 68/NA | 70/NA | 69      |
| Fontaine (1963)            | Mali            | Children in rural and urban primary schools                       | 746   | 1962   | 9–11   | 100 | 17.40 | 78    | 75    | –       |
| Heady (2003)               | Ghana           | Representative population sample                                  | 589   | 1988   | 9–11   | 82  | 15.80 | 72    | 71    | 62      |
| Jinabhai et al. (2004)     | South Africa    | Children from 11 rural primary schools in poor Vulamehlo district | 806   | ± 2002 | 8–10   | 100 | 13.9  | 72    | 68    | 67      |
| Kashala et al. (2005)      | Dem. Rep. Congo | Schoolchildren with ADHD and normal controls                      | 183   | 2003   | 8.5    | 100 | 15    | 79    | 74    | –       |
| Knoetze et al. (2005)      | South Africa    | Xhosa-speaking primary school students in peri-urban Eastern Cape | 172   | ± 2002 | 7.5–11 | 100 | 17.21 | 77    | 73    | –       |
| Okonji (1974)              | Nigeria         | Children in private school in Lagos                               | 73    | 1972   | 8–11   | 100 | 23.52 | 94    | 96    | –       |
| Ombredane et al. (1956)    | Dem. Rep. Congo | Children of “very underdeveloped” Asalampasu tribe                | 151   | ± 1955 | 6–11   | 79  | 14.50 | 76    | 76    | –       |
| Tzuriel & Kaufman (1999)   | Ethiopia        | Ethiopian Jews immigrated to Israel                               | 29    | ± 1992 | 6–7    | 100 | 15.60 | 94    | 92    | –       |
| Veii & Everatt (2005)      | Namibia         | Bilingual primary schoolchildren                                  | 116   | ± 2002 | 7–11   | 100 | 15.8  | 77    | 74    | –       |
| Wolff et al. (1995)        | Eritrea         | Orphans and refugee children during war                           | 148   | 1990   | 4–7    | NA  | 12.4  | 87    | 85    | –       |

Note: Edu: Percentage of sample formally educated; M: raw mean CPM score; IQ: IQ estimate; IQ FE: IQ estimate with correction for Flynn Effect; IQ Lynn: IQ estimate by Lynn (2006); and NA: not applicable.

formal schooling (defined as 3 years of education or less). These 734 uneducated test-takers had an *N*-weighted average IQ of approximately 71, which is considerably below the overall average. Note that the SPM may lack validity in test-takers without formal schooling (Dague, 1972), but lower scores among non-schooled test-takers may also reflect lower ability. One may want to correct for this underrepresentation of unschooled persons, as the percentage of unschooled young people is around 20%<sup>4</sup> in sub-Saharan Africa. A rough stratification for educational level can be achieved by adding 2900 fictional uneducated cases with an IQ of 71 to the total sample. This lowers the average IQ by one point to 77.2, which is almost equal to the estimate based on only the representative samples.

### 2.2.2. Coloured Progressive Matrices

Table 2 includes the results from the 16 samples in which the CPM was administered to African children (total *N* = 5466). The average IQs vary from 68 to 94. The *N*-weighted average of the twelve samples equals 78.1, (median 78, *SD* = 6.5). If we exclude Fahrmeier's (1975) study, in which the CPM was administered in a non-standard fashion, we arrive at an average IQ of 78.7 (median 78, *SD* = 6.1). Therefore, the average IQ of children in sub-Saharan Africa on the CPM appears to be 78 or 79. When compared to an average IQ of 100 for the US, this IQ among sub-Saharan African children equals 80 or 81. With a Flynn Effect correction of 2 IQ points per decade, these average IQs should be lowered by 2.5 points.

Besides the Fahrmeier sample, the samples that scored relatively lowly are the children from poor rural areas tested by Jinabhai et al. (2004), a sample of Ethiopian orphans (Aboud, Samuel, Hadera, & Addus, 1991), and the representative sample of Ghanaian children (Glewwe & Jacoby, 1992; Heady, 2003). The low average IQ of the orphans is perhaps due to the harsh circumstances that such children often encounter (but see Wolff, Tesfai, Egasso, & Aradom, 1995). Moreover, average IQ in rural areas is often lower than it is in urban areas (e.g., Loehlin, 2000). However, the low average IQ of the representative sample of children in Ghana is unexpected, given that Ghana is relatively well-developed (UN Development Programme, 2005). The low scores may be explained in part by the fact that the tests were administered in the children's houses. As the principle investigator put it: “[the test-takers] may have been sitting in a chair or even on the ground” while taking the tests (P. Glewwe, personal communication,

January, 17, 2006). This may have lowered the scores. Two recent representative standardization samples in Kenya (Costenbader & Ngari, 2001) and South Africa (Knoetze, Bass, & Steele, 2005) show average IQs of around 80.

Of the total of 5318 children of school-age, about 793 children (15%) did not attend school. This is slightly below the average population estimates of school attendance in current day sub-Saharan Africa. The number of rural children and urban children in the samples in Table 2 appears to roughly reflect the population distribution in sub-Saharan Africa. Finally, of the 16 samples considered, we consider four to be representative for a given population. Although definitive statements require completely stratified random population samples, the data in Table 2 provide a reasonable estimate of average IQ of African children on the CPM.

Because of a lack of sound UK norms for these groups, we did not consider several adult and adolescent samples that were administered the CPM. These samples were considered by Lynn (and Vanhanen) who accorded low mean IQs to these samples. It is certainly the case that the adult samples studied by Berry (1966) and others (Berlioz, 1955; Binnie Dawson, 1984; Kendall, 1976) showed low CPM averages as compared to western samples (Measso et al., 1993). In two studies with the CPM (Berry, 1966; Binnie Dawson, 1984), the authors deliberately sampled adults with very little knowledge of western culture. The representative samples of wage laborers from Nairobi and Dar es Salaam (Boissiere et al., 1985) scored roughly equal to a sample of adults from Italy and San Marino aged 40–59 (Measso et al., 1993), which does not lend much credibility to the low estimate of their mean IQ by Lynn (and Vanhanen). Even if we had included these adult and adolescent samples, average IQ based on the CPM scores in sub-Saharan Africa would not change very much. The reason is that most of these low scoring samples are small, whereas the carefully selected large sample (*N* = 1907) of adolescents in Uganda (Heyneman & Jamison, 1980) showed an average score equivalent to an average IQ above 79.<sup>5</sup>

<sup>5</sup> Lynn (2006) states that the 1907 primary school students tested with the CPM in Uganda (Heyneman & Jamison, 1980) are 11 years old, but most of these students are around 13. It is important to note that the score reported in Lynn's source (Heyneman & Jamison, 1980) is based on the correct number out of 33 instead of 36 items (Heyneman, 1975), but Lynn's source does not mention this. The first three items were used for instruction, so the average score needs to be raised by 3 points (these items typically have *p*-values of 1). If we add three points to the score and employ the CPM to SPM conversion (J. C. Raven et al., 1996), we can compare scores to the SPM norm table for the correct age range. This results in a rough estimate of an average IQ of 79. Due to the ceiling effect discussed earlier, this figure is likely to be too low.

<sup>4</sup> Based on UNESCO estimates of gross enrollment ratio in primary education over the period 1970–2003.

### 2.3. Conclusion on the average IQ of Africans

Our first aim was to estimate average IQ of Africans on the basis of published data of the CPM and SPM. The means of all 56 samples from Tables 1 and 2 are plotted against sample sizes in Fig. 1, which also contains the samples we excluded but that were considered by Lynn (2006). As can be seen, the largest samples are also those that are closest to the overall mean of all samples. The symmetry of the plot does not suggest appreciable publication bias. It is also clear that the exclusion of studies had little effect on our estimate and that the means vary widely between samples. Based on the SPM and CPM tests, we conclude that the average IQ in Africa lies somewhere around 78 (UK norms) or 80 (US norms). This estimate is 11 IQ points higher than the estimate of Lynn (and Vanhanen). There are some large discrepancies between our IQ estimates and those of Lynn (and Vanhanen). We refer to Wicherts (2007)<sup>6</sup> for a discussion of these discrepancies. From Tables 1 and 2, it is apparent that the samples not considered by Lynn (and Vanhanen) have considerably higher average IQ than the samples that they missed. Lynn (and Vanhanen) did not consider certain datasets despite the fact that these were present in the same sources from which they drew other datasets (Crawford Nutt, 1976; MacArthur et al., 1964; Raveau, Elster, & Lecoutre, 1976). They provided no reasons for this.

### 3. The Flynn Effect

In the western world, average IQs have shown remarkable gains over the course of the twentieth century (Flynn, 1984, 1987, 2007). These gains have been largest for non-verbal tests once considered relatively impervious to cultural influences, such as the Raven's (Brouwers et al., 2009). For instance, in The Netherlands an unaltered version of the SPM was administered to male military draftees from 1952 to 1982. The 1982 cohort scored approximately 20 IQ points higher than the 1952 cohort (Flynn, 1987). In this section, we consider whether a Flynn Effect has occurred among Africans on the Raven's tests.

#### 3.1. Flynn Effect in Africa

The results in Table 1 are based on data from diverse samples, varying in age and in nationality. Therefore, any secular trend in these data represents a tentative indication of African IQ trends. Nevertheless, the adult samples (ages 17 and higher) are fairly comparable with respect to age, because they all include young adults (even the Raveau samples only include a handful of cases above the age of 40). In the study of adult trends, we excluded the university samples, because these are all quite recent. We studied the Flynn Effect in the current samples by comparing all raw scores to the norms from the older standardization samples (i.e., 1938 for children and 1948 for adults). In the more recent samples this resulted in higher IQs than the values reported in Table 1. The results for the adult samples are plotted in Fig. 2. In this figure, we present four separate ( $N$ -weighted) regression lines of IQ on year of administration, namely for South Africa (based on 6 samples), Congo-Zaire (4 samples), and Zambia (3 samples), as well as for a combination of remaining countries (3 samples). All four lines show a clear upward trend over the years. The steepness of the regression lines suggests the presence a considerable Flynn Effect among African adults on the SPM. On average, these regression lines suggest increases of 7 IQ points per decade, which is comparable to that reported for male adults in the Netherlands from 1952–1982 (Flynn, 1987). Since a comparison of the adult samples from the different eras does not provide a compelling reason to believe that the samples are incomparable (e.g., in terms of educational attainment), the rise in adult samples appears to be a robust phenomenon.

<sup>6</sup> Jelte Wicherts' dissertation is available online at: <http://www.repository.naturalis.nl/document/44999>.

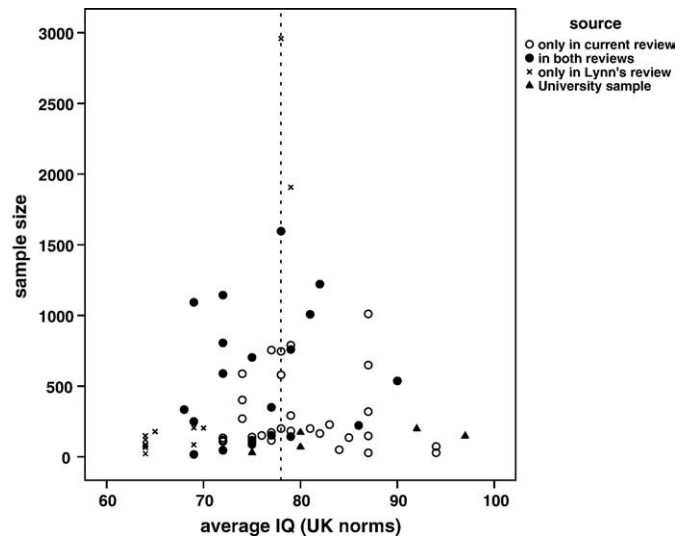


Fig. 1. Estimates of average IQ of Africans on the basis of SPM and CPM from our review and the review by Lynn (2006).

In Fig. 3, we present data concerning the SPM in children. We only depict the overall regression line for the Flynn Effect, because the lines of the individual countries vary greatly in slope and intercept. Overall, the Flynn Effect is significant ( $p < .001$ ) in adults, but not in children ( $p > .05$ ). There is no apparent rise in the children's samples, but a more definitive indication of a Flynn Effect among African children on the SPM should await samples that differ in time but are comparable in other respects.

Using the CPM, Daley et al. (2003) documented a Flynn Effect in two comparable samples of children from rural Kenya. If we exclude these two Kenyan samples, there is no clear indication of a Flynn Effect in the remaining samples in Table 2. The number of studies with the CPM in Africa is fairly small, and most recent samples are from rural areas. The various child samples are far from ideal to study the Flynn Effect. The absence of clear gains among children may be due to the fact that the older samples are primarily of school-going children in times when school attendance in Africa was mainly restricted to higher SES levels. Hence, differences in sampling, related to factors

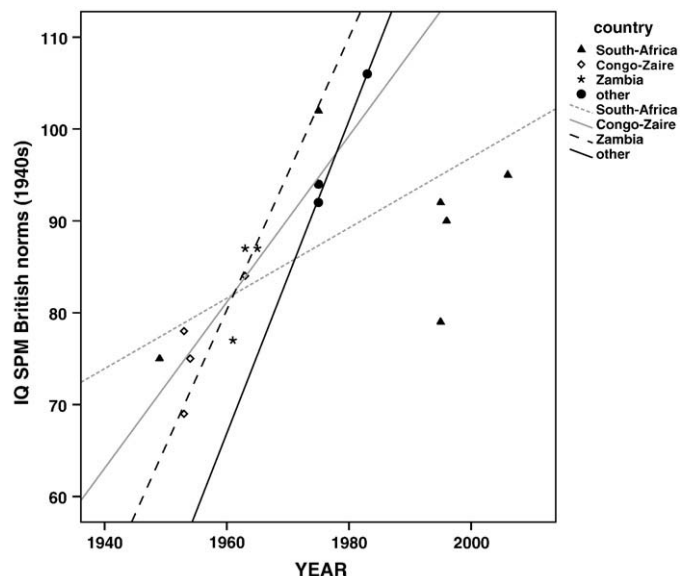


Fig. 2. Secular trends in IQ for samples of African children on the SPM.



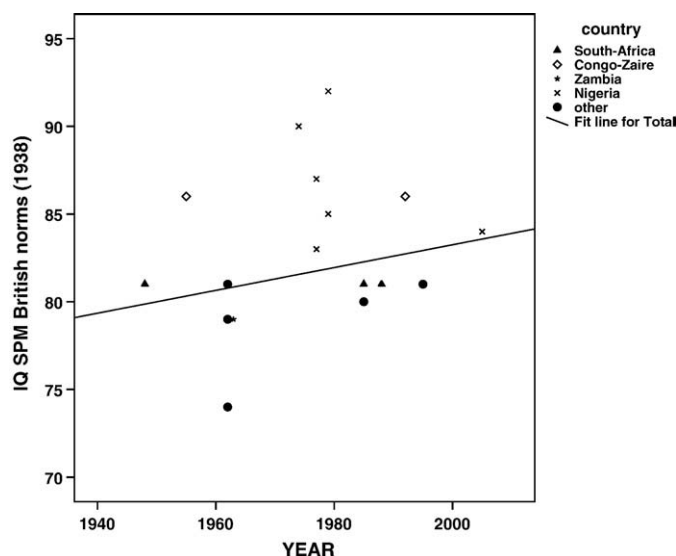


Fig. 3. Secular trends in IQ for samples of African adults on the SPM.

such as SES, schooling etc., may well be an important factor. To fully explore the issue of the Flynn Effect among African children on the both the CPM and the SPM, further research must employ samples that are comparable in terms of such demographic variables.

Proposed causes of the Flynn Effect include improvements in test-specific skills (Greenfield, 1998; Wicherts et al., 2004), improvements in nutrition (Lynn, 1989, 1990), urbanization (Barber, 2005), improvements in health care (Williams, 1998), a trend towards smaller families (Zajonc & Mullally, 1997), increases in educational attainment (Ceci, 1991), greater environmental complexity (Schooler, 1998), and the working of genotype by environment correlation in the increasing presence of more intelligent others (Dickens & Flynn, 2001). Many of these environmental variables have not undergone the improvement in developing sub-Saharan African countries that they have in the developed world over the last century. This suggests that the Flynn Effect has great potential in sub-Saharan Africa (Wicherts, Borsboom, & Dolan, 2010b).

#### 4. Measurement problems and psychometric comparability

A person's IQ and a person's level of latent general intelligence or *g* cannot simply be equated for the simple reason that IQ tests are fallible instruments. In Africa, IQ tests are often not administered in conditions comparable to those in developed countries. In fact, IQ test administration in Africa often occurs on the ground, on veranda's, under trees, or in overcrowded and sparsely furnished classrooms (e.g., Berry, 1983; Fahrmeier, 1975). Such non-standard test settings, combined perhaps with harsh climatic circumstances (cf. Sternberg, 2004), are likely to depress performance.

Moreover, the claim that the Raven's tests are "devoid of cultural content" (e.g., Templer & Arikawa, 2006, p. 122) does not sit well with the following measurement problems. Several items in the Raven's tests contain geometric shapes, which have no names in many African languages (Bakare, 1972). It is not uncommon in (rural) Africa to come across test-takers who are unfamiliar with color-printed material (Giordani, Boivin, Opel, Dia Nseyila, & Lauer, 1996), or who are inexperienced with using a pencil (Badri, 1965). Giving such test-takers, a paper-and-pencil test with unknown coloured geometric shapes (e.g., CPM) is not likely to produce test scores that accurately reflect *g*. Unfamiliarity with the stimulus material in western IQ tests (van den Briel et al., 2000) is only one of possible cultural factors that may affect performance of African test-takers when diagrammatic, non-verbal intelligence tests, such as the CPM or SPM, are used to assess general cognitive ability. For instance, responding to a multiple

choice test format may be entirely new to some African test-takers (Irvine, 1966). It cannot be assumed that a standard instruction engenders sufficient understanding of what is required (e.g., Kendall et al., 1988; MacArthur et al., 1964). Such problems have led some authors to conclude that intelligence testing and test-taking are strongly culturally determined (Berry, 1974, 1976; Greenfield, 1997; Irvine, 1969b; Nell, 2000; Sternberg, 2004). The degree to which this is so is unknown, but clearly measurement problems associated with IQ testing in Africa require careful consideration.

Others do claim that IQ test scores are comparable across cultures (Lynn, 2006; Rushton & Jensen, 2005). However, before one can interpret IQ test score differences across individuals or groups in terms of some latent cognitive ability (here, *g*), several conditions must be met. Necessary, but not sufficient, conditions for this interpretation concern the reliability and validity of tests. That is, the test scores must show some level of consistency, either internally, or in repeated testing. In addition, the test scores should correlate with other measures of cognitive ability (i.e., convergent validity). Factor analysis may be employed to elucidate the cognitive ability factors involved in test performance. A comparison of test scores across groups requires that the test scores are comparable in terms of the latent cognitive dimensions that a test is supposed to measure. In psychometrics, this condition is called measurement invariance. In the next section, we discuss these psychometric issues in turn. In Section 4.5 we consider the practical utility of the Raven's in educational settings.

##### 4.1. Reliability

Table 3 includes results from studies of the reliability of the Raven's tests in African samples. Reliabilities are generally above 0.80, which is comparable to those found in western samples (J. C. Raven et al., 1996). In Wober's (1969) sample, however, the retest reliability after six months was .59, which is low. The retest reliability in Verhaegen (1956) and Laroche (1959) was in the .50s after one year, which is again low. Interestingly, there exists a positive rank-order correlation between reliability and year of publication of the studies in Table 3 ( $r_s = .58$ ,  $p < .05$ , two-tailed). This indicates that reliability is greater in more recent samples. Nonetheless, the reliability of the Raven's tests appears to be sufficient in educated African samples. Verhaegen's study is the only study of the reliability of the Raven's tests among uneducated Africans and the results suggest that reliability in such samples may be poor.

##### 4.2. Convergent validity

The Raven's tests are considered good indicators of *g* in western samples (Carroll, 1993; Jensen, 1980). Consequently, Raven's test scores correlate highly with scores on many other cognitive ability tests. In his authoritative review, Burke (1958) found that the correlation between Raven's tests and (full scale) IQs from other batteries was around .64. Burke also reported the correlation between Raven's tests scores and non-verbal, verbal, memory, and spatial (sub) tests in 33 western samples. The convergent validities in these studies varied from .26 to .80. We fitted a random effects meta-analytic model (on the Fischer-transformed correlation coefficients ( $Z_r$ ) and with ( $N-3$ ) as weight) on Burke's correlations, which resulted in a mean  $Z_r$  of .559 (95% confidence interval: .497–.622; Dersimonian-Laird-estimated variance component .022), which equates to  $r = .507$ . Burke's results on the basis of western studies can serve as a benchmark for the convergent validity in African samples.

In numerous studies the Raven's tests were administered to Africans alongside other cognitive ability tests. The correlations of the SPM and CPM with other cognitive tests in 23 African samples are reported in Table 4. This table includes all the studies in our review that reported correlation coefficients in African samples. The table

**Table 3**  
Reliability results for the Progressive Matrices in Africa.

| Source                     | Country      | Sample description                       | N    | Test | Type     | Rel. |
|----------------------------|--------------|--|------|------|----------|------|
| Costenbader & Ngari (2001) | Kenya        | Representative sample of school children | 50   | CPM  | rt.2 wk. | .84  |
|                            |              |  | 1370 | CPM  | Alpha    | .87  |
| Crawford Nutt (1976)       | S-Africa     | Children from high school in Soweto      | 228  | SPM  | KR-20    | .82  |
| Hoorweg (1976)             | Uganda       | Malnourished adolescents                 | 60   | CPM  | KR-20    | .87  |
| Jahoda (1956)              | Ghana        | Urban school boys                        | 317  | SPM  | rt.2 wk. | .82  |
| Kendall (1976)             | South Africa | Male recruits in mining industry         | 149  | CPM  | KR-20    | .87  |
| Laroche (1959)             | Congo–Zaire  | Boys in schools in Elizabethville        | 222  | SPM  | rt.1 wk. | .84  |
|                            |              |  | 73   | SPM  | rt.1 yr. | .56  |
|                            |              |  | 76   | SPM  | rt.2 yr. | .63  |
| Maqsdud (1997)             | South Africa | Standard 9 students from Mmabatho        | 30   | SPM  | rt.1 mo. | .83  |
| Nkaya et al. (1994)        | Congo–Braz.  | Secondary school children                | 88   | SPM  | rt.2 wk. | .91  |
| Owen (1992)                | South Africa | Children from secondary schools          | 1093 | SPM  | Alpha    | .93  |
| Pons (1974)                | Zambia       | Bemba males employed in mining           | 152  | SPM  | Retest   | .82  |
| Pons (1974)                | Zambia       | Bemba males employed in mining           | 1011 | SPM  | Retest   | .88  |
| Rushton & Skuy (2000)      | South Africa | University students in psychology        | 173  | SPM  | Alpha    | .91  |
| Rushton et al. (2002)      | South Africa | University students in engineering       | 198  | SPM  | Alpha    | .87  |
| Sternberg et al. (2001)    | Kenya        | Children from rural area                 | 85   | CPM  | Alpha    | .85  |
| Verhaegen, (1956)          | Congo–Zaire  | Unschooling adults from Kasai            | 67   | SPM  | rt.1 wk. | .88  |
|                            |              |  |      | SPM  | rt.1 yr. | .55  |
| Wober (1969)               | Nigeria      | Male factory workers                     | 86   | SPM  | rt.6 mo. | .59  |

Note: Alpha: Cronbach's Alpha; KR-20: Kuder–Richardson's formula 20; retest: retest reliability, time span unknown; rt.1 yr: retest reliability after 1 yr; and rt.1 wk.: retest reliability after 1 wk.

**Table 4**  
Convergent validity of SPM and CPM in African samples.

| Source                                | Country      | Sample description  | Setting | N   | Test | Other test (s)  | Correlations       |
|---------------------------------------|--------------|---|---------|-----|------|---|--------------------|
| Aboud et al. (1991)                   | Ethiopia     | Children in an orphanage in Jimma                                 | U       | 134 | CPM  | Conservation measure  | .43                |
| Berry, (1966)                         | Sierra Leone | Temne adults  | R       | 122 | CPM  | Koh's blocks, and embedded figures  | .43, and .17       |
| van den Briel et al. (2000)           | Benin        | Schoolchildren in four rural villages                             | R       | 196 | CPM  | Block design, closure, and concentration  | .05, .08, and .18  |
|                                       |              |   |         |     |      | Exclusion, fluency, mazes, and K-ABC hand movement                                      | -.01, and .14      |
| Crawford Nutt (1976)                  | South Africa | Children from high school in Soweto                               | U       | 102 | SPM  | Dominoes, symco, and numerical ability  | .59, .19 and .53   |
| Crawford Nutt (1976)                  | South Africa | Children from high school in Soweto                               | U       | 102 | SPM  | Blox, vocabulary, and mental alertness  | .40, .17, and .51  |
| (Grieve & Viljoen, 2000; Sonke, 2001) | South Africa | Impoverished University students in rural Venda                   | UNI     | 29  | SPM  | Dominoes, symco, and numerical ability  | .43, .01, and .48  |
| Hoorweg (1976)                        | Uganda       | Malnourished and control adolescents                              | U/R     | 90  | CPM  | Austin maze   | .25                |
|                                       |              |   |         |     |      | Halstead–Reitan category test   | .40                |
| Irvine (1963)                         | Zimbabwe     | Standard 6 students from Harare                                   | U       | 204 | SPM  | Vocabulary, arithmetic, and Port. maze  | .29, .58, and .41  |
|                                       |              |   |         |     |      | Block design, and memory for design   | .66, and .47       |
| Jinabhai et al. (2004)                | South Africa | Children from 11 rural primary schools in poor Vulamehlo district | R       | 806 | CPM  | Knox cubes, learning, and Inc. learning   | .43, .30, and .13  |
|                                       |              |   |         |     |      | Mental alertness, and compound series   | .56, and .49       |
| Kendall (1976)                        | South Africa | Male recruits in mining industry                                  | U/R     | 149 | CPM  | Number ability, and spiral nines  | .52, and .46       |
|                                       |              |   |         |     |      | Auditory verbal learning Test, symbol digit modalities test, and group mathematics test | .10–.36            |
| Lynn & Holmshaw (1990)                | South Africa | Children from state primary schools                               | U/R     | 350 | SPM  | Sorting I, sorting II, and cube construction  | .46, .44, and .52  |
|                                       |              |   |         |     |      | Tripod assembly, pattern repr., and circles   | .38, .50, and .57  |
| MacArthur et al. (1964)               | Zambia       | Representative sample of students in class 6                      | U/R     | 684 | SPM  | Form series, form perc., and fret repetition  | .52, .38, and .41  |
|                                       |              |   |         |     |      | Fret continuation   | .52                |
| MacArthur et al. (1964)               | Zambia       | Representative sample of students in Form II                      | U/R     | 442 | SPM  | Choice RT, odd-man-out RT (both reverse scored)   | -.01, and -.05     |
|                                       |              |   |         |     |      | Spiral nines, and comprehension   | .51, and .37       |
| MacArthur et al. (1964)               | Zambia       | Technical college students  | UNI     | 61  | SPM  | Vocabulary, spelling, computation   | .29, .15, and .30  |
|                                       |              |   |         |     |      | Boxes, figures, mental alertness, vocabulary, spelling, and computation                 | .32, .38, and .46  |
| Morakinyo (1985)                      | Nigeria      | Psychiatric out-patients  | U       | 16  | SPM  | Boxes, figures, mental alertness, comprehension, vocabulary, spelling and computation   | .08, .40, and .52  |
|                                       |              |   |         |     |      | Verbal learning   | .59                |
| Notcutt (1950)                        | South Africa | Zulus in primary school is near Durban                            | U/R     | 108 | SPM  | Span of apprehension, and cube pattern  | .21, and .52       |
|                                       |              |   |         |     |      | Zulu vocabulary   | .41                |
| Okonji (1974)                         | Nigeria      | Children in private school in Lagos                               | U       | 73  | CPM  | Auditory–visual pattern test  | .11                |
|                                       |              |   |         |     |      | Advanced Progressive Matrices   | .60                |
| Rushton et al. (2003)                 | South Africa | University students   | UNI     | 187 | SPM  | Mental alertness, and perc. flex. English. lang   | .34, .36, and .12  |
|                                       |              |   |         |     |      | Verb. reason, reading comp. and Sci. info.  | .26, .22, and .11  |
| Silvey (1972)                         | Uganda       | Buganda secondary school children                                 | U/R     | 136 | SPM  | Current affairs, cler. speed, and graph read.   | .08, .21, and .11  |
|                                       |              |   |         |     |      | Mental alertness, and English language  | .07, and .02       |
| Silvey (1972)                         | Uganda       | Secondary school children in west-Uganda                          | U/R     | 73  | SPM  | Mental alertness, and English language  | .07, and .02       |
| Sternberg et al. (2001)               | Kenya        | Children from rural area  | R       | 85  | CPM  | Tacit knowledge, Mill-Hill, and Dhuolo Voc.   | -.16, .36, and .30 |
| Tzurriel & Kaufman (1999)             | Ethiopia     | Ethiopian Jews immigrated to Israel                               | R       | 29  | CPM  | Analogical thinking   | .49                |
|                                       |              |   |         |     |      | Inferential Thinking  | .19                |
| Wober (1969)                          | Nigeria      | Male factory workers  | U       | 86  | SPM  | Embedded figures test   | .48                |

Note: setting U: urban; R: rural; R/U: Both rural and urban; and UNI: University.



includes the type of test used and an indicator of the setting of the study, i.e., Urban, Rural, or mixed Urban/Rural.

Some studies support the convergent validity of the SPM and CPM among Africans. The highest correlation was found in a study at the University of Witwatersrand in South Africa, where the SPM scores correlated .60 with the scores on SPM's more difficult counterpart, the Advanced Progressive Matrices (APM). In a study among 85 adolescents in Kenya, CPM correlated reasonably well with two vocabulary tests, but non-significantly with a test of practical intelligence (Sternberg et al., 2001). Moreover, SPM scores correlated considerably with a verbal learning task among healthy and unhealthy adults in Nigeria (Morakinyo, 1985), and with a perceptual learning potential test administered to 43 South African university students (Skuy et al., 2002). The latter data are not included in Table 4, because the precise correlation was not given. For this same reason, several additional studies were not included (Glewwe & Jacoby, 1992; Heyneman & Jamison, 1980; Kashala, Elgen, Sommerfeldt, Tylleskar, & Lundervold, 2005; Munroe, Munroe, & Brasher, 1985; Veii & Everatt, 2005).

The results of several studies raise serious questions about the convergent validity of the Raven's tests among Africans. In contrast to results in Britain, Japan, and Hong Kong (cf. Lynn, 1991), Lynn and Holmshaw (1990) did not find the SPM scores in their sample of Black South African children ( $N=350$ ) to correlate significantly with the cognitive components of reaction time tasks. Similarly, correlations of the CPM with other cognitive ability tests were generally below .40 in two other studies (Jinabhai et al., 2004; Okonji, 1974). In a study of 196 children in rural Benin, the CPM correlated quite lowly with seven other cognitive tests (van den Briel et al., 2000).

We fitted a random effects meta-analytic model to the correlations in Table 4, with ( $N-3$ ) as weight. In studies that included more than one correlation, we used the mean correlation. The mean meta-analytic correlation was  $Zr = .328$  (95% confidence interval: .250–.406), which equates to  $r = .317$ . There was a clear indication of heterogeneity of correlations:  $Q = 123$ ,  $DF = 22$ ,  $p < .001$ . Dersimonian–Laird estimate of variance was equal to .027. We found that the convergent validity in African samples was significantly lower than the validity results in the 33 western samples from Burke's (1958) literature review:  $Z = 4.60$ ,  $p < .001$ . As is the case with Burke's results, the results in Table 4 are based on a mixture of homogeneous and heterogeneous samples of school children, university students, and adults. Hence, there is no reason to expect that restriction of range differs greatly between the western and the African studies. Fig. 4 depicts the transformed corre-

lations against the inverse of the standard error for both the African samples and the western samples in Burke's study. The difference in convergent validity of the SPM/CPM between African and western samples is quite large. The symmetry of both funnel plots does not suggest any publication bias, nor does a formal test of publication bias (Sterne & Egger, 2005).

In a mixed model meta-regression of the statistics in Table 4, we found that the Urban/Rural variable significantly predicted the validity coefficients ( $Z = 2.69$ ,  $p < .01$ ). The average predicted  $Zr$  was .189 ( $r = .187$ ) in the rural samples and .443 ( $r = .416$ ) in urban samples, with the coefficient in the mixed Rural/Urban samples in between (predicted  $Zr = .316$ ,  $r = .306$ ). Accordingly, we conclude that the convergent validity of the Raven's tests markedly lower in rural and mixed rural/urban African samples than in western samples. It is noteworthy that none of the studies in Table 4 involved uneducated test-takers. The validity of the Raven's tests among uneducated test-takers remains to be studied, but is likely to be poor (Dague, 1972), particularly in light of the poor retest reliability found by Verhaegen (1956).

#### 4.3. Factor analytical results

In factor analyses with other cognitive ability tests, the Raven's tests usually have large  $g$  loadings (i.e., correlation with  $g$ ; Carroll, 1993). Jensen (1980, 1998) noted that the  $g$  loadings of Raven's tests are normally around .80, and that these tests measure  $g$  and little else. This means that the SPM and CPM usually do not load appreciably on factors besides  $g$ . We now consider the factor analytic data that bear on the  $g$  loading and factorial purity of the SPM and CPM in Africa.

MacArthur (1973) factor analyzed data from two samples of Zambian teenagers who completed a large number of other cognitive ability tests including the SPM. His results showed that the five item sets of the SPM loaded on two different factors in one sample ( $N = 65$ ) and on three different factors in another sample ( $N = 65$ ). This suggests that the SPM is neither unidimensional nor factorially pure in these samples. Vernon (1950) reported on a factor analysis of scores on 13 tests observed in a large number of African recruits. He found that the scores on a revised version of the SPM showed an unrotated loading of .57 on the first factor, as well as a sizeable cross-loading of .34 on a second factor he interpreted in terms of physical factors and manual dexterity. Claeys (1972) factor analyzed data from 101 first-grade teachers from Kinshasa, who completed a battery of 24 tests including the SPM, and found that the commonality of the SPM was rather low (i.e., .38).

In order to study the  $g$  loading of the SPM and CPM in African samples, we ran exploratory factor analyses on the correlation matrices in the ten African samples in which the CPM or SPM was administered with at least four other cognitive ability tests (cf. Table 4). Our choice of the number of factors (which varied from 2 to 3) was based on the most parsimonious model that passed the exact fit test ( $p > .05$ ). We used VARIMAX rotation. Table 5 includes the summary results (detailed results are available on request). The CPM showed a clear cross-loading in Hoorweg's data of Ugandan teenagers, which is consistent with Hoorweg's (1976) own analyses. In MacArthur's sample of technical college students, the  $g$  loading of the SPM was .53, but the SPM also showed a sizeable loading of .35 on a spatial factor. This is in line with Irvine's (1969a) factor analytic study of these data. Likewise, in Kendall's sample the SPM loaded on three factors, and in one of the two samples reported by Crawford Nutt, SPM again showed a cross-loading on a factor other than  $g$ .

The only two African samples in which the SPM lived up to expectations were MacArthur's sample of Form II students and one of the two secondary school samples from Crawford Nutt's study. In these samples, the SPM did not show a cross-loading (cf. Irvine, 1969a). In five of the ten samples, the SPM and CPM showed cross loadings on non- $g$  factors, and in the sample of children from rural

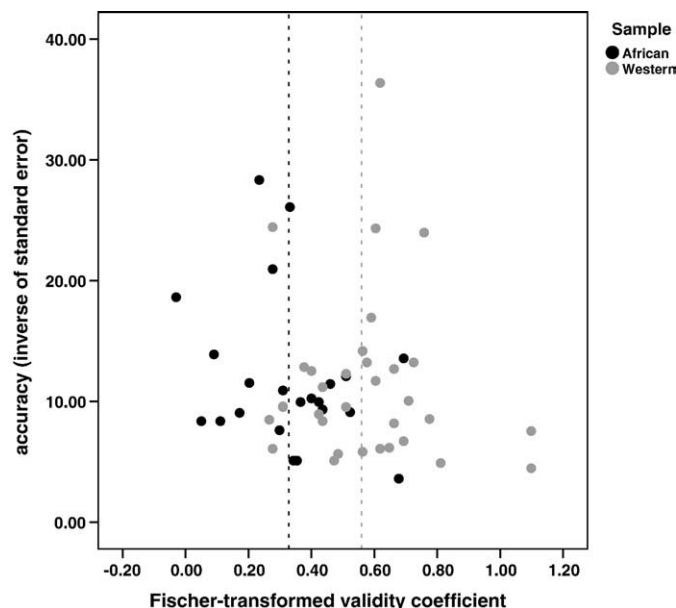


Fig. 4. Funnel plot of convergent validities in western and African samples.

**Table 5**  
Factor analysis results of SPM and CPM in African samples.

| Source                      | Country      | Sample description                           | N   | Test | $h^2$ | No of tests | <i>g</i> Loading | Loading 2nd fac. | Loading 3rd fac. |
|-----------------------------|--------------|--|-----|------|-------|-------------|------------------|------------------|------------------|
| van den Briel et al. (2000) | Benin        | Schoolchildren in four rural villages        | 196 | CPM  | .04   | 8           | .11              | .16              | –                |
| Crawford Nutt (1976)        | South Africa | Children from high school in Soweto          | 102 | SPM  | .50   | 5           | <b>.59</b>       | <b>.38</b>       | –                |
| Crawford Nutt (1976)        | South Africa | Children from high school in Soweto          | 102 | SPM  | .56   | 5           | <b>.74</b>       | .10              | –                |
| Hoorweg (1976)              | Uganda       | Malnourished and control adolescents         | 90  | CPM  | .62   | 9           | <b>.70</b>       | <b>.36</b>       | –                |
| Irvine (1963)               | Zimbabwe     | Standard 6 students from Harare              | 204 | SPM  | .49   | 5           | <b>.61</b>       | <b>.35</b>       | –                |
| Kendall (1976)              | South Africa | Male recruits in mining industry             | 149 | CPM  | .43   | 11          | <b>.43</b>       | <b>.27</b>       | <b>.42</b>       |
| MacArthur et al. (1964)     | Zambia       | Representative sample of students in class 6 | 684 | SPM  | .39   | 6           | <b>.62</b>       | .02              | –                |
| MacArthur et al. (1964)     | Zambia       | Representative sample of students in Form II | 442 | SPM  | .46   | 8           | <b>.65</b>       | .03              | .18              |
| MacArthur et al. (1964)     | Zambia       | Technical college students                   | 61  | SPM  | .41   | 8           | <b>.53</b>       | <b>.35</b>       | –                |
| Silvey (1972)               | Uganda       | Buganda secondary school children            | 136 | SPM  | .34   | 9           | <b>.56</b>       | .14              | .09              |

Note: based on ML-exploratory factor analysis and varimax rotation;  $h^2$ : communality; loadings larger than .25 in bold.

Benin, the CPM completely failed to load on the dominant factor. The average *g* loading in the nine samples was .55, which is considerably smaller than the *g* loading of the Raven's in western samples (Jensen, 1998). Combined, the results of these factor analyses do not support the assertion that the CPM and SPM are highly *g*-loaded tests among African test-takers. Moreover, it appears that in many African samples, the Raven's tests reflect other factors, in addition to *g*.

#### 4.4. Measurement invariance

Good validity and reliability within cultural groups are necessary, but not sufficient, requirements for understanding group differences in test performance. An additional necessary requirement is measurement invariance. Measurement invariance across groups implies that the relation between test scores and latent traits, which are supposed to underlie those scores, is identical across groups.<sup>7</sup> Measurement invariance is a starting point to understand the nature of group differences in test scores, so it is central to the question of the meaning of IQ of Africans. Measurement invariance can be tested by employing a measurement model in which this relation is explicitly modeled (Holland & Wainer, 1993; Meredith, 1993; Millsap & Everson, 1993). The relation between test scores and latent traits is central to the question of cross-cultural comparability of IQ test scores (e.g., Little, 1997; Poortinga & van de Vijver, 1988). The lack of measurement invariance suggests measurement bias, or, Differential Item Functioning (DIF), as it is called in Item Response Theory (IRT) modeling. DIF is said to be absent if, in a sufficiently restrictive measurement model (e.g., an unidimensional item response model), measurement parameters linking ability to tests scores are equal across groups.

We came across one recent study that tested for DIF of the SPM in an African sample (Taylor, 2008). In this study, the SPM item scores from 200 Black South African job applicants were tested for DIF against a sample of 178 White South African job applicants. In total 4 of the 60 items evidenced DIF, which the author deemed a small enough number to accept measurement invariance. To our knowledge, this is the only study that has used the methods of contemporary IRT to study DIF of the Raven's in Africa. Other studies used older methods.

Ombredane et al., in a study of the item characteristics in their samples (Ombredane, 1957; Ombredane, Robaye, & Plumail, 1956), found that the CPM showed a relatively large number of Guttman errors (e.g., Meijer & Sijtsma, 2001), which means that responses to relatively easy items included more errors than those to more difficult items. Irvine

(1969b) conducted a factor analysis of SPM items, and concluded that, unlike in the western samples he studied, the SPM was not unidimensional in African samples. Combined with the factor analytic results at the scale level, it is therefore quite likely that the Raven's tests measure other factors, in addition to *g*, in African samples.

In a series of studies, Rushton et al. (Rushton, 2002; Rushton & Skuy, 2000; Rushton, Skuy, & Bons, 2004; Rushton et al., 2002; Rushton, Skuy, & Fridjhon, 2003) studied whether the SPM and the APM have similar item characteristics for Whites and Blacks in South Africa (cf. Owen, 1992). Rushton et al. (2004) claim that these studies establish the construct validity for IQ tests among Africans. However, all but one study (Owen, 1992) involved university students, and a perusal of Tables 3, 4, and 5 indicates that the reliability and convergent validity are relatively good among university samples. It is unclear whether the results from university samples generalize to lower-scoring samples. More importantly, in none of the studies by Owen and Rushton et al. was DIF studied across groups. Instead, two simple methods were employed to study bias of Raven's items. The first method is based on the rank-order correlation between item *p*-values across groups. Rushton et al. and Owen reported correlations in the .90s between item *p*-values between western and African samples. Earlier, Irvine (1966) had reported a correlation of only .695 between the item *p*-values in British school children and *p*-values in a representative sample of 1600 pupils from then Rhodesia. However, it is important to note that this method (Thurstone, 1925), and various refinements (e.g., the Delta-Plot method; Angoff & Ford, 1972) have been criticized extensively in the psychometric literature for not being sensitive to DIF (Lord, 1980; Shepard, Camilli, & Williams, 1985).

Rushton and colleagues used a second method in studying the comparability of Raven's scores between African and western samples (Rushton, 2002; Rushton, Bons, Vernon, & Cvorovic, 2007; Rushton & Skuy, 2000; Rushton et al., 2004; Rushton et al., 2002, 2003). This method involves correlating the vector of group differences in item's *p*-values with the vector of item-total correlations (i.e., point-biserial or biserial item-total correlations). Wicherts and Johnson (2009) have shown that this method does not address the issue of measurement invariance. Finally, we were unable to find any studies of measurement invariance of cognitive test batteries that included the Raven's tests.

In sum, although one study supported measurement invariance (Taylor, 2008), other studies paint a different picture (Irvine, 1966, 1969a; Ombredane, 1957; Ombredane et al., 1956). Hence, there is little empirical support for measurement invariance of the Raven's tests between African and western samples.

#### 4.5. Criterion validity in educational settings

A lack of measurement invariance renders cross-cultural comparisons of Raven's test scores problematic, but does not necessarily mean that the these tests have no practical utility in professional and educational settings within African countries. In western samples, scores on the SPM and CPM predict test-takers' performance in school

<sup>7</sup> Mellenbergh (1989) provided a general definition of measurement invariance that is expressed in terms of the conditional distribution of manifest test scores *Y*, denoted by  $f(Y|N)$ . According to this definition, measurement invariance with respect to *v* holds if:  $f(Y|N, v) = f(Y|N)$ , where *N* denotes the scores on the latent variable (i.e., latent ability) underlying the manifest random variable *Y* (i.e., the measured variable), and *v* is a grouping variable, which defines the nature of groups (e.g., nationality). One important implication of this definition is that if measurement invariance holds, the expected test score of a person with a certain latent ability (i.e., *N*) is independent of group membership.

settings reasonably well. Raven et al. (1996) indicated that the SPM correlates between .20 and .60 with school grades, examination results, or teacher ratings. Burke (1958) reported correlations between SPM and student performance or grade point average from .14 to .39 in six western samples. We found 16 African samples in which the predictive or criterion validity of the Raven's tests in educational settings could be determined (additional studies did not report raw correlations so could not be used; e.g., Daley et al., 2003; Fahrmeier, 1975). These results are reported in Table 6, with a description of the sample and criteria. We again considered potential differences between urban and rural samples in predictive and criterion validity, but found no systematic differences.

Several studies supported the predictive validity of Raven's tests. For instance, Zaiman et al. (2001) found that the SPM correlated reasonably well with college performance, and Maqsd (1980a) reported significant correlations between the SPM and school grades. In contrast, some studies failed to find support for the criterion validity SPM in Africa (e.g., Laroche, 1959; Ogunlade, 1978). However, all correlations in Table 6 are positive and a meta-analytic model fitted on the data indicated that the average meta-analytic  $Z_r$  was .271 [.191–.351] (which corresponds to  $r = .265$ ). There was some heterogeneity in correlations:  $Q = 75.9$ ,  $DF = 15$ ,  $p < .01$ , with a Dersimonian–Laird-estimate of variance of .019. These results are not markedly different from results in western samples. To conclude, there is some support for the predictive utility of the Raven's tests in educational settings in Africa. However, given the relatively weak correlations between the Raven's and educational outcomes, additional, more specific, cognitive tests may improve prediction and so may be necessary for particular purposes (e.g., selection).

#### 4.6. Conclusion on psychometric properties

The empirical support for the validity of the SPM and CPM as measures of  $g$  among Africans is mixed. Compared to results from western studies, the convergent validity of the Raven's tests among Africans is poor, especially in rural samples. An average correlation of .20 or .30 between the Raven's tests and other cognitive ability tests suggests that the Raven's tests in Africa are not as strongly  $g$ -loaded as they are in western samples. Indeed, factorial studies suggests that the Raven's test scores of Africans show relatively weak  $g$  loadings, and often show cross-loading on factors other than  $g$ . At the same time, the criterion validity of these tests among Africans in educational settings is comparable to western samples (see also Kendall et al., 1988). However, with an average criterion validity coefficient of .265, the practical utility of the Raven's tests remains quite weak. In light of the many claims concerning the unsuitableness of the Raven's tests in Africa (Irvine, 1969b; Ogunlade, 1978; Verhaegen, 1956; Wober, 1966, 1975), and the

relatively low scores in some samples, more research into the construct validity of the Raven's tests among Africans is needed.

More importantly, we found only one study that employed modern psychometric techniques to study DIF of Raven's items between non-African and African samples. We are not aware of any study that studied measurement invariance of the Raven's in a confirmatory factor analysis framework. Thus, it is unclear whether Raven's tests afford an adequate comparison of western and African samples in terms of the construct of  $g$ .

## 5. General discussion

Our review of the literature on the performance of Africans on the Raven's tests showed that the average IQ of Africans on the Raven's tests is lower than the average IQ in western countries. However, the average IQ of Africans is not as low as Lynn (and Vanhanen) and Malloy (2008) maintained. The majority of studies on IQ test performance of Africans not taken into account by Lynn (and Vanhanen) and Malloy showed considerably higher average IQs than the studies that they did review. We judge the reviews of Lynn (and Vanhanen) and Malloy to be unsystematic. These authors missed a large part of the literature on IQ testing in Africa, failed to explicate their inclusion and exclusion criteria, and made downward errors in the conversion of raw scores to IQs (Wicherts, 2007). Lynn (and Vanhanen)'s estimate of average IQ of Africans of around 67 is untenable. Our review indicates that it is about 78 (UK norms) or 80 (US norms). These means are somewhat lower than the means of Africans on other IQ tests, which lie around 82 (Wicherts et al., 2010). These results undermine evolutionary theories of race differences in intelligence of Lynn (2006), Rushton (2000), and Kanazawa (2004) (Wicherts, Borsboom, & Dolan, 2010a; Wicherts et al., 2010b).

A given IQ test score should not be equated uncritically with a particular level of  $g$ , because IQ tests are fallible instruments, particularly for test-takers who are relatively unfamiliar with western culture as reflected in these IQ tests. IQ testing in Africa is a complicated affair (MacArthur et al., 1964; Nell, 2000). Based on our reading of the literature, validity studies of the SPM or CPM in Africa provide little support that these tests furnish adequate assessments of  $g$ . Although these tests predict academic outcomes to some degree, the correlations between Raven's test and other cognitive tests are generally lower in African samples than in western samples. Factor analyses show that the  $g$  loading of the Raven's tests is considerably smaller in African than in western samples. Our factor analyses indicate that  $g$  explains (on average) around 30% of the variance of Raven's test scores among Africans. In the majority of factor analytic studies of African's performance on the Raven's tests, the results suggest that these tests load on other factors, in addition to  $g$ . Hence,

**Table 6**

Criterion validity of SPM and CPM in African samples.

| Source                  | Country      | Sample description                                 | N   | Test | Criterion                               | $r$ |
|-------------------------|--------------|--|-----|------|---|-----|
| Aboud et al. (1991)     | Ethiopia     | Children in an orphanage in Jimma                  | 134 | CPM  | School grades (with age partialled out) | .36 |
| Crawford Nutt (1977)    | South Africa | Children from high school in Soweto                | 102 | SPM  | School grades                           | .14 |
| Crawford Nutt (1977)    | South Africa | Children from high school in Soweto                | 102 | SPM  | School grades                           | .24 |
| Durojaiye (1973)        | Uganda       | Secondary school students                          | 540 | SPM  | School performance                      | .14 |
| Irvine (1963)           | Zimbabwe     | Standard 6 students from Harare                    | 204 | SPM  | School exams                            | .46 |
| Laroche (1959)          | Congo-Zaire  | Boys in schools in Elizabethville                  | 218 | SPM  | School grades                           | .17 |
| MacArthur et al. (1964) | Zambia       | Representative sample of Class 6 students          | 684 | SPM  | Total school grades                     | .44 |
| MacArthur et al. (1964) | Zambia       | Representative sample of Form II students          | 442 | SPM  | Total school grades                     | .27 |
| MacArthur et al. (1964) | Zambia       | Mine farm youth                                    | 85  | SPM  | Manual training performance             | .23 |
| Maqsd (1980a)           | Nigeria      | Boys from modern primary schools                   | 60  | SPM  | English and Arithmetic                  | .57 |
| Maqsd (1980a)           | Nigeria      | Boys from traditional primary schools              | 60  | SPM  | English and Arithmetic                  | .29 |
| Ogunlade (1978)         | Nigeria      | Secondary school students in towns in West-Nigeria | 537 | SPM  | School grades                           | .15 |
| Silvey (1972)           | Uganda       | Secondary school children in west-Uganda           | 73  | SPM  | School certificate and school exam      | .06 |
| Silvey (1972)           | Uganda       | Buganda secondary school children                  | 136 | SPM  | School certificate and school exam      | .11 |
| Sternberg et al. (2001) | Kenya        | Children from rural area                           | 85  | CPM  | English & Math achievement              | .19 |
| Zaiman et al. (2001)    | South Africa | Disadvantaged university students                  | 147 | SPM  | Final grade score                       | .33 |



there is a clear need for more work on the factorial structure of these tests in African samples (Irvine, 1969a). Our psychometric findings do not sit well with the notion that the Raven's tests are culturally fair, so we would recommend caution in the use of these tests with Africans for selection purposes in education and the global labor market.

Although one study (Taylor, 2008) supported measurement invariance of SPM items in an African and a non-African sample using an established psychometric method, such methods have been rarely applied to study the psychometric meaning of Africans' scores on the Raven's tests or other IQ tests (Wicherts et al., 2010). Most studies that used alternative methods to study psychometric comparability of Raven's items showed mixed or inconclusive results. Without rigorous studies of measurement invariance, it remains unknown whether the low average performance on the SPM and CPM of Africans reflect *g*, measurement bias, test-specific abilities (e.g., test sophistication), lower-level abilities (e.g., spatial abilities), or a combination of such effects. The degree to which non-*g* effects may have led to an underestimation of ability in sub-Saharan African samples remains unknown. Given the very low scores in some African samples, we consider such an underestimation to be plausible. It is also noteworthy that several authors found that retesting may improve the validity of the Raven's tests (Laroche, 1959; Wober, 1969), so it would be interesting to consider the validity of scores from retests in future work (see also Skuy et al., 2002).

The true meaning of IQ differences between western and African samples can be addressed in carefully designed cross-cultural studies that employ appropriate psychometric modeling. What is required is a study in which testing conditions across groups are controlled, and in which it is ascertained that test instructions are clear to all test-takers. Dynamic assessment approaches (e.g., Carlson & Wiedl, 1979; Lidz & Elliott, 2000) may be useful in this regard. This study should involve a battery of tests, each of which can be studied for DIF. Then, one could establish by using psychometric modeling whether between-group mean differences are on the (higher order) latent factor called *g* (Dolan, 2000; Dolan & Hamaker, 2001; Lubke, Dolan, & Kelderman, 2001; Lubke, Dolan, Kelderman, & Mellenbergh, 2003). Suppose we could establish that tests are fully measurement invariant, and that between-group differences are mainly (or entirely) due to between-group differences in *g*. This opens the door to address the question why groups differ with respect to the latent variable that we call *g*. Only after we have tackled the measurement problem, can we start to study the causes of group difference in *g*, rather than group differences with respect to sum scores on tests, whose meaning remains ambiguous.

Although the implications of our psychometric findings for the potential of the Flynn Effect in sub-Saharan Africa remain unclear, the Raven's tests and other IQ tests have shown robust increases in many populations (Daley et al., 2003; Flynn, 2007). So suppose that there were a well-validated IQ test that showed measurement invariant scores between westerners and Africans. Even then, lower IQs of Africans still would not support Lynn and Vanhanen's (2002, 2006) assertion that countries in sub-Saharan Africa are poorly developed economically because of their low "national IQ". Wicherts, Borsboom, and Dolan (2010b) found that "national IQs" are rather strongly confounded with the developmental status of countries. Given the well-documented Flynn Effect, we know that "national IQs" are subject to change. An average IQ around 80 among Africans may appear to be low, but from a historical perspective this average is not low at all. A representative sample of British adults, who took the SPM in 1948 would have an average IQ of 81 in terms of the British norms of 1992 (J. C. Raven, 1960; J. C. Raven et al., 1996). Using older British norms, the average IQ of Africans would be much closer to 100. This is evident in Fig. 2, where we compared SPM scores of Africans to older norms. In this figure, the average IQ of several African samples is near or above 100.

Present-day sub-Saharan Africa is one of the poorest regions in the world and the home to some of the world's most deprived children.

The majority of sub-Saharan children are chronically malnourished, not only from lack of food but particularly from food lacking vital elements related to both physical growth and intellectual development. It has been estimated that up to 70% of rural children live in absolute poverty and 90% suffer severe deprivation (Gordon, Nancy, Pantazis, Pemberton, & Townsend, 2003). A substantial number of sub-Saharan African children are under-educated. According to Garcia, Gillian, and Dunkelberg (2008), only about 12% of sub-Saharan African children have attended preschool, and this generally for well less than a year. They note that children who do not attend or have only minimal experience in pre-primary school tend to do less well in primary school than children who have had that experience. Further, it is important that the preschool experience be successful. For example, Jaramillo and Mingat (2008) have shown that children who have a poor experience in preschool and have to repeat a year or part of a year have a high drop-out rate in primary school ( $r = -0.875$ ). The probability of preschool without repetition and who complete primary school is low but positive ( $r = 0.209$ ). With or without preschool experience, approximately only fifty-five% of 10–14 year-olds in sub-Saharan Africa complete primary school.

It is widely agreed that education, or the lack of it, is among the most critical problems in developing countries and regions. Both the length and quality of schooling are major factors contributing to personal development and intellectual growth (Blair, Gamson, Thorne, & Baker, 2005; Cahen & Cohen, 1989; Cliffordson & Gustafsson, 2008), and a region's economic development (Card, 1999; Hanushek & Kimko, 2000). For developing nations, investment in primary and secondary education is particularly important and appears to be cost effective (Psacharopoulos & Patrinos, 2002).

Many of the variables that have been proposed as causes of the Flynn Effect (e.g., Barber, 2005; Ceci, 1991; Dickens & Flynn, 2001; Greenfield, 1998; Lynn, 1989; Schooler, 1998; Williams, 1998) have yet to undergo improvement in developing sub-Saharan African countries that they have enjoyed in the developed world over the last century. Because the environmental variables that potentially contribute to enhanced IQ have yet to improve in the countries of sub-Saharan Africa, we regard the Flynn Effect as still in its infancy. There is in fact a considerable empirical support that (mal)nutrition (Grantham-McGregor & Baker-Henningham, 2007; Sigman & Whaley, 1998), health (Williams, 1998), sanitation (Boivin et al., 1993), and schooling (Ceci, 1991) have an effect on IQ. The UN have included such variables in the so-called Millennium Goals, i.e., they are targeted for improvement by 2015 (United Nations, 2005). The formulation of the Millennium Goals provides an interesting opportunity to evaluate the effect of these factors on IQ levels in sub-Saharan Africa. There is now a clear indication that the Flynn Effect seems to have come to a halt in developed nations (Flynn, 2007). It is, therefore, reasonable to think that as circumstances in sub-Saharan Africa improve, the IQ gap between western samples and African samples will diminish.

The future will tell whether average IQ among sub-Saharan Africans will show gains similar to those found in western countries. However, those who claim that IQ among Africans is low because of genetic or evolutionary factors (Kanazawa, 2008; Lynn, 2006; Rushton, 2000; Templer & Arikawa, 2006), may wish to take into consideration the fact that African countries are underdeveloped (Wicherts et al., 2010b). On the basis of our review, we cannot preclude the possibility that genetic factors may play a role in the low IQ levels among Africans. However, to the best of our knowledge, heritability of IQ within African populations has not been the subject of behavior genetic study. Given that heritability can vary as a function of the quality of environmental factors (e.g., SES see Turkheimer, Haley, Waldron, D'Onofrio, & Gottesman, 2003), it is quite possible that heritability of IQ in African populations is lower than it is in the west. With respect to the role of genetic factors between populations (African vs. western), it is very difficult to progress beyond speculation in the light of ecological correlations, which do not necessarily support

causal interpretations. However, in view of the established potential of the Flynn Effect, the collinearity among mean IQs and a variety of socio-economic indices, and the developmental status of many African countries, we consider environmental factors to be of primary importance in the interpretation of current IQ levels among Africans.

The vast literature on IQ testing with the Raven's tests in Africa does not support James Watson's pessimism concerning the prospects of Africa. It is true that Africans show lower average IQs as compared to contemporary western norms, although the IQ gap is substantially smaller than Lynn (and Vanhanen) have maintained. More importantly, there is little scientific basis for the assertion that the observed lower IQs of Africans are evidence of lower levels of general in-

telligence or *g*. The validity of the Raven's tests among Africans needs to be studied further before these tests can be used to assess Africans' cognitive ability in educational and professional settings. There are several reasons to expect increases in IQ levels among sub-Saharan Africans in the coming decades.

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### Appendix A. Converting raw scores to IQs

For the SPM scores collected during the 1950 s and 1960 s, we used British norms of 1938 for children and 1940/1948 norms for adults (both of which are given in: J. C. Raven, 1960). For the data collected after 1965 (i.e., the midpoint of various SPM/CPM standardizations), we used British SPM norms of the 1979 standardization for children (J. Raven, 2000) and the 1992 standardization for adults (J. C. Raven et al., 1996). Likewise, we used the 1956 British norms (J. C. Raven, 1956) of the CPM for samples in the 1950 s and the first half of the 1960 s. For samples from 1966 onwards, we employed the British CPM norms of 1982 (J. C. Raven et al., 1990). We provide IQ estimates with and without a correction for the Flynn Effect. There are two reasons that a correction for the Flynn Effect may be unwarranted. First, the secular increase has not been documented in most African countries, and so we cannot reasonably correct for an effect, about which little is known in Africa. Second, we chose the year of 1965 as a midpoint, so the upward and downward corrections would be approximately balanced. Nevertheless, we checked whether the Flynn Effect correction makes a difference. The correction we used is based on Lynn (and Vanhanen)'s approach in which British norms on the SPM and CPM should be lowered by 2 IQ points per decade when data are more recent than the UK standardization, or 2 points per decade upwards when the data are older than the standardization (Lynn & Hampson, 1986).

Notwithstanding its perceived status as a good indicator of *g* (Jensen, 1998), it is important to note that Raven originally did not intend the Progressive Matrices test to be used as an IQ test. There are indeed several reasons for not converting Raven's scores to IQs. First, the Progressive Matrices are limited to a single test format. The test taker's unfamiliarity with this format may depress test performance. Second, in comparison to IQ batteries, such as the Wechsler scales (Wechsler, 1974, 1997), the number of items is relatively small in both the SPM (i.e., 60 items) and the CPM (i.e., 36 items). An additional problem arises in the translation from SPM/CPM raw scores to IQs in the extreme score ranges, where values in the norm tables show large leaps for particular age groups.

In our calculation of average IQs based on raw SPM and CPM scores, we exercised the upmost care. We did all calculations twice to minimize computational errors. Because most of these norms tables do not give percentiles for all raw scores, some inter- and extrapolation was necessary to arrive at percentile scores. In the few cases where a raw mean of a sub-sample was below the 1st percentile, we assigned an IQ of 64 (similar to the approach employed by Lynn & Vanhanen, 2002). In studies, where scores were reported for sub-samples, we first estimated IQs for the sub-samples, and then computed an *N*-weighted average of IQs for the entire sample. Whenever scores were given for a particular age range (e.g., 7–8 years), the average IQ was compared to the norms for the corresponding age groups (e.g., 7, 7½, 8, and 8½ year-olds). The average IQ was an average of these age-norms after weighing for sample size. If a data source indicated that the age distribution was peaked about a certain age, we adjusted our estimates accordingly. Our approach almost always resulted in an IQ estimate that was equal to the estimate that was based on the overall average raw score of the sample, when compared to the norm that corresponded to the average age of the entire sample. If not, we took the average IQ of both approaches. From one source (Fontaine, 1963), we had to estimate the average age of child samples on the basis of school grade levels. Where medians were given, we took the median as an estimate of the mean. In one source (Morakinyo, 1985), percentile scores were reported, so we transformed these to approximate raw means in referring the scores to a more recent norm table. All steps in the estimation of IQs are available upon request to the first author. We note that in determining IQ, it is conceivable that the aggregation of raw scores from different test-takers with varying ages does not necessarily match the average of IQs, when these are computed at the individual level. Both the first and the second author independently estimated percentile scores of 21 of the sub-samples that took either the SPM or the CPM. The inter-rater reliability of these sub-samples turned out to be 0.92. The inter-rater reliability of the nine IQ estimates on the basis of these sub-samples was 0.96, which is sufficiently high for the current purposes. Nonetheless, the assignment of IQ values for the SPM and the CPM remains problematic, and the values we provide are only given in order to arrive at a rough estimate of average IQ.

### Appendix B

An overview of excluded studies using the SPM or the CPM in Africa.

| Source                    | Country      | Sample                               | Reason for exclusion  | Test | N    | IQ Lynn         |
|---------------------------|--------------|--------------------------------------|---|------|------|-----------------|
| Ahmed (1989)              | Sudan        | School children from Khartoum        | Khartoum is not part of sub-Saharan Africa                        | SPM  | 100  | 72              |
| Bellis et al. (1988)      | Mali         | Children and adults with cretinism   | Inappropriate sample, no appropriate CPM norms                    | CPM  | 413  | –               |
| Berlioz (1955)            | Cameroon     | Specialized workmen                  | No appropriate CPM norms for adults                               | CPM  | 80   | 64 <sup>1</sup> |
| Berry (1966)              | Sierra Leone | Temne adults                         | No appropriate CPM norms for adults                               | CPM  | 122  | 64 <sup>1</sup> |
| Binnie Dawson (1984)      | Sierra Leone | Temne agriculturalists               | No appropriate CPM norms for adults                               | CPM  | 22   | 64 <sup>1</sup> |
| Boissiere et al. (1985)   | Kenya        | Wage laborers                        | No appropriate CPM norms for adults                               | CPM  | 205  | 69 <sup>1</sup> |
| Boissiere et al. (1985)   | Tanzania     | Wage laborers                        | No appropriate CPM norms for adults                               | CPM  | 179  | 65 <sup>1</sup> |
| Erwee (1981)              | S-Africa     | University students                  | Test not administered in its entirety                             | SPM  | 110  | –               |
| Heyneman & Jamison (1980) | Uganda       | Primary school students              | No appropriate CPM norms for this age group                       | CPM  | 1907 | 73 <sup>2</sup> |
| Hoorweg (1976)            | Uganda       | Malnourished and control adolescents | No appropriate CPM norms for this age group, inappropriate sample | CPM  | 90   | –               |

## Appendix B (continued)

| Source                             | Country      | Sample  | Reason for exclusion                                 | Test | N    | IQ Lynn         |
|------------------------------------|--------------|---|--|------|------|-----------------|
| Irvine (1963)                      | Zimbabwe     | Standard 6 children in Harare                   | Time limit imposed, test coaching                    | SPM  | 217  | –               |
| Jahoda (1956)                      | Ghana        | Urban school boys                               | Time limit imposed                                   | SPM  | 317  | –               |
| (Jansen, Richter, & Griesel, 1992) | South Africa | Street children living in a shelter             | No appropriate CPM norms for this age group          | CPM  | 44   | –               |
| Jukes et al. (2006)                | Gambia       | Teenagers from villages where malaria is common | No appropriate CPM norms for this age group          | CPM  | 579  | –               |
|                                    |              |   | Test adapted   |      |      |                 |
| Kendall (1976)                     | Mozambique   | Male novices                                    | No appropriate CPM norms for adults                  | CPM  | 149  | 64 <sup>1</sup> |
| Klingelhofer (1967)                | Tanzania     | Secondary school students                       | Time limit imposed to prevent comparison             | SPM  | 2959 | 78              |
| Maqsud (2003)                      | S-Africa     | Standard 7 pupils                               | Group selected because of low scores on SPM          | SPM  | 40   | –               |
| Morgaut (1959)                     | Various      | Adult workers                                   | SPM test not administered in its entirety            | SPM  | 5863 | –               |
| Osuji (1985)                       | Nigeria      | Physically disabled adults                      | Inappropriate sample                                 | SPM  | 330  | –               |
| Pillay (2003)                      | S-Africa     | Mentally retarded children                      | Inappropriate sample, raw mean score lacking         | CPM  | 50   | –               |
| Silvey (1972)                      | Uganda       | Secondary school children                       | Time limit imposed                                   | SPM  | 209  | –               |
| Sternberg et al. (2001)            | Kenya        | Children from rural area                        | No appropriate CPM norms for this age group          | CPM  | 85   | 69              |
| Verhaegen (1956)                   | Congo-Zaire  | Unschooling adults from Kasai                   | Test format was "rather confusing" to test-takers    | SPM  | 67   | 64              |
| Wober (1969)                       | Nigeria      | Male factory workers                            | Reliability and validity insufficient                | SPM  | 86   | 64              |
| Zindi (1994)                       | Zimbabwe     | School children                                 | Test not fully administered, raw mean score lacking. | SPM  | 204  | 70              |

Notes. <sup>1</sup>On the basis of a study in Italy and San Marino (Measso et al., 1993), underestimation of adult IQ with the CPM is approximately 25 IQ points; <sup>2</sup>average IQ should be approximately 79, cf. Footnote 5.

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